Selected Solutions for Chapter 11: Hash Tables

Solution to Exercise 11.2-1

For each pair of keys k, l, where $k \neq l$, define the indicator random variable $X_{kl} = I\{h(k) = h(l)\}$. Since we assume simple uniform hashing, $\Pr\{X_{kl} = 1\} = \Pr\{h(k) = h(l)\} = 1/m$, and so $E[X_{kl}] = 1/m$.

Now define the random variable Y to be the total number of collisions, so that $Y = \sum_{k \neq l} X_{kl}$. The expected number of collisions is

$$E[Y] = E\left[\sum_{k \neq l} X_{kl}\right]$$

= $\sum_{k \neq l} E[X_{kl}]$ (linearity of expectation)
= $\binom{n}{2} \frac{1}{m}$
= $\frac{n(n-1)}{2} \cdot \frac{1}{m}$
= $\frac{n(n-1)}{2m}$.

Solution to Exercise 11.2-4

The flag in each slot will indicate whether the slot is free.

- A free slot is in the free list, a doubly linked list of all free slots in the table. The slot thus contains two pointers.
- A used slot contains an element and a pointer (possibly NIL) to the next element that hashes to this slot. (Of course, that pointer points to another slot in the table.)

Operations

Insertion:

- If the element hashes to a free slot, just remove the slot from the free list and store the element there (with a NIL pointer). The free list must be doubly linked in order for this deletion to run in O(1) time.
- If the element hashes to a used slot *j*, check whether the element *x* already there "belongs" there (its key also hashes to slot *j*).
 - If so, add the new element to the chain of elements in this slot. To do so, allocate a free slot (e.g., take the head of the free list) for the new element and put this new slot at the head of the list pointed to by the hashed-to slot (*j*).
 - If not, *E* is part of another slot's chain. Move it to a new slot by allocating one from the free list, copying the old slot's (*j*'s) contents (element *x* and pointer) to the new slot, and updating the pointer in the slot that pointed to *j* to point to the new slot. Then insert the new element in the now-empty slot as usual.

To update the pointer to j, it is necessary to find it by searching the chain of elements starting in the slot x hashes to.

- *Deletion:* Let *j* be the slot the element *x* to be deleted hashes to.
 - If x is the only element in j (j doesn't point to any other entries), just free the slot, returning it to the head of the free list.
 - If x is in j but there's a pointer to a chain of other elements, move the first pointed-to entry to slot j and free the slot it was in.
 - If x is found by following a pointer from j, just free x's slot and splice it out of the chain (i.e., update the slot that pointed to x to point to x's successor).
- *Searching:* Check the slot the key hashes to, and if that is not the desired element, follow the chain of pointers from the slot.

All the operations take expected O(1) times for the same reason they do with the version in the book: The expected time to search the chains is $O(1 + \alpha)$ regardless of where the chains are stored, and the fact that all the elements are stored in the table means that $\alpha \leq 1$. If the free list were singly linked, then operations that involved removing an arbitrary slot from the free list would not run in O(1) time.

Solution to Problem 11-2

a. A particular key is hashed to a particular slot with probability 1/n. Suppose we select a specific set of k keys. The probability that these k keys are inserted into the slot in question and that all other keys are inserted elsewhere is

$$\left(\frac{1}{n}\right)^k \left(1 - \frac{1}{n}\right)^{n-k}$$

Since there are $\binom{n}{k}$ ways to choose our k keys, we get

$$Q_k = \left(\frac{1}{n}\right)^k \left(1 - \frac{1}{n}\right)^{n-k} \binom{n}{k}.$$

- **b.** For i = 1, 2, ..., n, let X_i be a random variable denoting the number of keys that hash to slot *i*, and let A_i be the event that $X_i = k$, i.e., that exactly *k* keys hash to slot *i*. From part (a), we have $Pr \{A\} = Q_k$. Then,
 - $P_{k} = \Pr \{M = k\}$ $= \Pr \{\left(\max_{1 \le i \le n} X_{i}\right) = k\}$ $= \Pr \{\text{there exists } i \text{ such that } X_{i} = k \text{ and that } X_{i} \le k \text{ for } i = 1, 2, ..., n\}$ $\leq \Pr \{\text{there exists } i \text{ such that } X_{i} = k\}$ $= \Pr \{A_{1} \cup A_{2} \cup \cdots \cup A_{n}\}$ $\leq \Pr \{A_{1}\} + \Pr \{A_{2}\} + \cdots + \Pr \{A_{n}\} \quad \text{(by inequality (C.19))}$ $= nQ_{k}.$
- *c*. We start by showing two facts. First, 1 1/n < 1, which implies $(1 1/n)^{n-k} < 1$. Second, $n!/(n-k)! = n \cdot (n-1) \cdot (n-2) \cdots (n-k+1) < n^k$. Using these facts, along with the simplification $k! > (k/e)^k$ of equation (3.18), we have

$$Q_{k} = \left(\frac{1}{n}\right)^{k} \left(1 - \frac{1}{n}\right)^{n-k} \frac{n!}{k!(n-k)!}$$

$$< \frac{n!}{n^{k}k!(n-k)!} \qquad ((1 - 1/n)^{n-k} < 1)$$

$$< \frac{1}{k!} \qquad (n!/(n-k)! < n^{k})$$

$$< \frac{e^{k}}{k^{k}} \qquad (k! > (k/e)^{k}) .$$

d. Notice that when n = 2, $\lg \lg n = 0$, so to be precise, we need to assume that $n \ge 3$.

In part (c), we showed that $Q_k < e^k/k^k$ for any k; in particular, this inequality holds for k_0 . Thus, it suffices to show that $e^{k_0}/k_0^{k_0} < 1/n^3$ or, equivalently, that $n^3 < k_0^{k_0}/e^{k_0}$.

Taking logarithms of both sides gives an equivalent condition:

$$3 \lg n < k_0 (\lg k_0 - \lg e)$$

=
$$\frac{c \lg n}{\lg \lg n} (\lg c + \lg \lg n - \lg \lg \lg n - \lg e).$$

Dividing both sides by $\lg n$ gives the condition

$$3 < \frac{c}{\lg \lg n} (\lg c + \lg \lg n - \lg \lg \lg n - \lg e)$$
$$= c \left(1 + \frac{\lg c - \lg e}{\lg \lg n} - \frac{\lg \lg \lg n}{\lg \lg n} \right).$$

Let *x* be the last expression in parentheses:

$$x = \left(1 + \frac{\lg c - \lg e}{\lg \lg n} - \frac{\lg \lg \lg n}{\lg \lg n}\right)$$

We need to show that there exists a constant c > 1 such that 3 < cx.

Noting that $\lim_{n\to\infty} x = 1$, we see that there exists n_0 such that $x \ge 1/2$ for all $n \ge n_0$. Thus, any constant c > 6 works for $n \ge n_0$.

We handle smaller values of n—in particular, $3 \le n < n_0$ —as follows. Since n is constrained to be an integer, there are a finite number of n in the range $3 \le n < n_0$. We can evaluate the expression x for each such value of n and determine a value of c for which 3 < cx for all values of n. The final value of c that we use is the larger of

- 6, which works for all $n \ge n_0$, and
- $\max_{3 \le n < n_0} \{c : 3 < cx\}$, i.e., the largest value of c that we chose for the range $3 \le n < n_0$.

Thus, we have shown that $Q_{k_0} < 1/n^3$, as desired.

To see that $P_k < 1/n^2$ for $k \ge k_0$, we observe that by part (b), $P_k \le nQ_k$ for all k. Choosing $k = k_0$ gives $P_{k_0} \le nQ_{k_0} < n \cdot (1/n^3) = 1/n^2$. For $k > k_0$, we will show that we can pick the constant c such that $Q_k < 1/n^3$ for all $k \ge k_0$, and thus conclude that $P_k < 1/n^2$ for all $k \ge k_0$.

To pick *c* as required, we let *c* be large enough that $k_0 > 3 > e$. Then e/k < 1 for all $k \ge k_0$, and so e^k/k^k decreases as *k* increases. Thus,

$$Q_k < e^{\kappa}/k^{\kappa}$$

$$\leq e^{k_0}/k^{k_0}$$

$$< 1/n^3$$

for $k \geq k_0$.

e. The expectation of M is

$$E[M] = \sum_{k=0}^{n} k \cdot \Pr\{M = k\}$$

$$= \sum_{k=0}^{k_0} k \cdot \Pr\{M = k\} + \sum_{k=k_0+1}^{n} k \cdot \Pr\{M = k\}$$

$$\leq \sum_{k=0}^{k_0} k_0 \cdot \Pr\{M = k\} + \sum_{k=k_0+1}^{n} n \cdot \Pr\{M = k\}$$

$$\leq k_0 \sum_{k=0}^{k_0} \Pr\{M = k\} + n \sum_{k=k_0+1}^{n} \Pr\{M = k\}$$

$$= k_0 \cdot \Pr\{M \le k_0\} + n \cdot \Pr\{M > k_0\},$$

which is what we needed to show, since $k_0 = c \lg n / \lg \lg n$.

To show that $E[M] = O(\lg n / \lg \lg n)$, note that $Pr\{M \le k_0\} \le 1$ and

$$Pr \{M > k_0\} = \sum_{k=k_0+1}^{n} Pr \{M = k\}$$

=
$$\sum_{k=k_0+1}^{n} P_k$$

<
$$\sum_{k=k_0+1}^{n} 1/n^2 \qquad (by part (d))$$

<
$$n \cdot (1/n^2)$$

=
$$1/n .$$

We conclude that

$$E[M] \leq k_0 \cdot 1 + n \cdot (1/n)$$

= $k_0 + 1$
= $O(\lg n / \lg \lg n)$.